

For Raegly Distribution Simulation with the Help of Kendall Distribution Function Archimedean Copula Parameter Estimation

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Abstract

In this paper, we model the dependence structure between random variables that we generated dependent Raeghly distribution using Archimedean copula and Kendall distribution function. In connection with this, we define basic properties of copulas and their nonparametric method. The aim of Kendall distribution function is selected suitable copula function for using data set. For dependence structures of the data set, we calculated Kendall Tau and Spearman Rho values which are nonparametric. Based on this method, parameters of copula are obtained. To explain the relationship between the variables, three Archimedean copula families were used; Gumbel, Clayton, Frank, Ali Mikhail Haq and Joe copula. Nonparametric estimation of copula parameters and we find the suitable Archimedean copula family for this data set.

Key Words: Copula functions; Kendall Tau; Kendall Distribution Function; Raeghly distribution.

MSC 2010: 62A01, 62H20, 62H12.

1. Introduction

Copulas were introduced to first in the context of theory metric spaces. The statistical properties and applications of copulas has been developing in recent years. In 1959 A.Sklar introduced the general notions of a copula (1981 By.B.Schweizer and E.F.Wolff) [6]. A copula function is links univariate marginal to their multivariate distribution. We can model relationship between random variables. Copula function is presenting the dependence structure and it is providing degree of dependence structure. Copula is continuous

transformation and invariant under increasing. Copulas can use for modeling dependence in several applied fields such as econometric, finance and actuarial studies. This article explores Archimedean copula definitions us to reduce the study of multivariate copula to a single univariate function. Throughout the paper we work bivariate Archimedean copulas; Clayton, Gumbel, Frank, Ali Mikhail Haq and Joe.

2. Material and Method

The copula is defined as a function $C: [0,1]^2 \rightarrow [0,1]$ that ensures the boundary conditions

- ✓ $C(u, 0) = C(0, u) = 0$ and $C(u, 1) = C(1, u) = u, \forall u \in [0,1]$
- ✓ $\forall (u_1, u_2, v_1, v_2) \in [0,1]^4$, such that $u_1 \leq u_2, v_1 \leq v_2$,
- ✓ $C(u_2, v_2) - C(u_2, v_1) - C(u_1, v_2) + C(u_1, v_1) \geq 0$.

Ultimately, for twice differentiable and 2-increasing property can be replaced by the condition

$$c(u, v) = \frac{\partial^2}{\partial u \partial v} C(u, v) \geq 0 \quad (2.1)$$

Where $c(u, v)$ is the copula density. In the following for n uniform random U_1, U_2, \dots, U_n variables the joint distribution function C is defined

$$C(u_1, u_2, \dots, u_n, \theta) = P(U_1 \leq u_1, U_2 \leq u_2, \dots, U_n \leq u_n)$$

Here θ is dependence parameter. [1,2,3,4,5,6,7,8,9,10,11,12,14,15,16,17,18,19,20]

2.1. Sklar Theorem

Let X and Y be random variables with continuous distribution functions F_X and F_Y , with $F_X(X)$ and $F_Y(Y)$ uniformly distributed on the interval $[0,1]$. Then there exists a copula such that for all x, y in \mathbb{R} ,

$$F_{XY}(X, Y) = C(F_X(X), F_Y(Y)) \quad (2.2)$$

The copula C for (X, Y) is just the joint distribution function for the couple $F_X(X), F_Y(Y)$ provided F_X and F_Y continuous [1,2,3,4,5,6,7,8,9,10,11,12,14,15,16,17,18,19,20].

2.2. Archimedean Copula

Let φ define a function $\varphi : [0,1] \rightarrow [0, \infty]$ that is continuous and supplies:

- ✓ $\varphi(1) = 0, \varphi(0) = \infty$.
- ✓ For all $t \in (0,1), \varphi'(t) < 0, \varphi$ is decreasing, for all $t \in (0,1) \varphi''(t) \geq 0, \varphi$ is convex.

φ has an inverse $\varphi^{-1} : [0, \infty] \rightarrow [0,1]$, that is to say this equation has the similar properties out of $\varphi^{(-1)}(0) = 1$ and $\varphi^{(-1)}(\infty) = 0$. The Archimedean copula is given by

$$C(u, v) = \varphi^{(-1)}[\varphi(u) + \varphi(v)]. \quad (2.3)$$

[1,2,3,4,5,6,7,8,9,10,11,12,14,15,16,17,18,19,20].

2.3. Clayton copula

The Archimedean copula is defines with the help of generator function $\varphi(t) = \frac{t^{-\theta} - 1}{\theta}$

$$C_\theta(u, v) = (u^{-\theta} + v^{-\theta} - 1)^{1/\theta} \quad (2.4)$$

θ is the copula parameter restricted to $(0, \infty)$ [13].

2.4. Frank copula

The Archimedean copula is defines with the help of generator function;

$$\begin{aligned} \phi(t) &= -\ln \frac{e^{-\theta t} - 1}{e^{-\theta} - 1}, \theta \in \mathbb{R} / \{0\}; \\ C_\theta(u, v) &= -\frac{1}{\theta} \ln \left(1 + \frac{(e^{-\theta u} - 1)(e^{-\theta v} - 1)}{(e^{-\theta} - 1)} \right) \end{aligned} \quad (2.5)$$

θ is the copula parameter restricted to $(0, \infty)$ [13].

2.5. Gumbel copula

The Archimedean copula is defines with the help of generator function $\phi(t) = (-\ln t)^\theta, \theta \geq 1$;

$$C_\theta(u, v) = \exp\left(-[(-\ln u)^\theta + (-\ln v)^\theta]^{1/\theta}\right) \quad (2.6)$$

θ is the copula parameter restricted to $[1, \infty)$ [13].

2.6. Ali Mikhail Haq Copula

This Archimedean copula is defined with the help of generator function $\varphi(t) = \ln[1 - \theta(1-t)]/t$

$$C_\theta(u, v) = \frac{uv}{1 - \theta(1-u)(1-v)} \quad (2.7)$$

Where θ is the copula parameter restricted to $[-1,1]$ [13].

2.7. Joe Copula

This Archimedean copula is defined with the help of generator function $\varphi(t) = -\ln[1 - (1-t)^\theta]$

$$C_\theta(u, v) = 1 - \left[(1-u)^\theta + (1-v)^\theta - ((1-u)^\theta (1-v)^\theta) \right]^{1/\theta} \quad (2.8)$$

Where θ is the copula parameter restricted to $[1, \infty)$ [13].

2.7. The Nonparametric Estimation

The Archimedean Copula submit each copula has statement that connections its parameters to associated Kendal Tau and Spearman Rho. In this study only the relationships contain Kendal Tau that given in table.

Table1. The link between Archimedean copulas and Kendall Tau.

Family	Range of θ	τ
Gumbel	$\theta \in [1, \infty)$	$\frac{\theta-1}{\theta}$
Clayton	$\theta \in [0, \infty)$	$\frac{\theta}{\theta+2}$
Frank	$\theta \in (-\infty, \infty)$	$1 - \frac{4}{\theta} [1 - D_1(\theta)]$
Ali Mikhail Haq	$\theta \in [-1, 1]$	$\frac{3\theta-2}{3\theta} - \frac{2(1-\theta)^2 \ln(1-\theta)}{3\theta^2}$
Joe	$\theta \in [1, \infty)$	$1 + \frac{4}{\theta} D_1(\theta)$

Here D is debye functions [5,6,15]

2.8. Kendall Distribution Function and Properties

In the past, it proposed a nonparametric method for forecasting the dependence function of a double of random variables for Archimedean copula (Genest and Rivest, 1993). The state of emphasizing a probability model for independent observations $(x_1, y_1), \dots, (x_n, y_n)$ from a bivariate non Gaussian distribution function $H(X, Y)$ might be reduced by denoted H and its marginal of F_X and F_Y , its related dependence function C . C is the association copula with generator ϕ and Kendall Distribution function the function given by

$$K(u) = \Pr\{C(U_1, \dots, U_n \leq u)\}$$

(Genest and Rivest, 1993) give that if C is Archimedean copula, forecast of Archimedean copula is singly defined by function on the space $(0,1)$;

$$K(u) = u - \frac{\phi(u)}{\phi'(u)}$$

a nonparametric estimation of K is shown by

$$K_n(u) = \sum_{j=1}^n I\{U_j \leq u\} / n + 1$$

To define the generator function ϕ , we show the paces; to forecast Kendall Tau value utilizing the non-parametric estimation and nonparametric forecast of K . For $K_n(u)$ nonparametric estimation of $K(u)$

- i) The nonparametric forecast of Archimedean copula Kendall Tau correlation coefficient using
- ii) Define the pseudo-observations

$$U_i = F_n(X_i, Y_i) =$$

$$\sum_{j=1}^n I\left[\left\{X_j \leq X_i, Y_j \leq Y_i\right\}\right] / n + 1, i = 1, 2, \dots, n$$

$$K_n(u) = \frac{(U_i \leq u)}{n + 1} = \frac{\text{number of } U_i \leq u}{n + 1}$$

- iii) Form a parametric estimation of K

$$K(u) = u - \frac{\phi(u)}{\phi'(u)}$$

- iv) The election of Archimedean copula that suitable for the data may be done by minimum a range

$$\int \left[K_{\phi_n}(u) - K_n(u) \right]^2 dK_n(u) \tag{2.7}$$

Schweitzer and Wolff, (1981), Cherubini and Luciano, (2001), Frees and Valdes, (1998), Genest et al., (2009).

Table 2. Kendall Distribution functions of Archimedean copulas.

Family	Generator $\phi(u)$	Generator first derivative $\phi'(u)$	The distribution function $K(u) = u - \frac{\phi(u)}{\phi'(u)}$
Gumbel	$(-\ln(u))^\theta$	$-\theta(\ln u)^{\theta-1} \frac{1}{u}$	$u - \frac{(u \ln u)}{\theta}$
Clayton	$u^{-\theta} - 1$	$-\theta u u^{-\theta-1}$	$u - \frac{(u^{\theta+1} - u)}{\theta}$
Frank	$-\ln\left(\frac{e^{-\theta u} - 1}{e^{-\theta} - 1}\right)$	$\frac{\theta}{1 - e^{\theta u}}$	$u - \frac{\ln \frac{e^{-\theta u} - 1}{e^{-\theta} - 1}}{\theta} (e^{-\theta u} - 1)$
Ali Mikhail Haq	$\ln[1 - \theta(1-u)]/u$	$\frac{\theta u - \ln[1 - \theta(1-u)][1 - \theta(1-u)]}{u^2[1 - \theta(1-u)]}$	$u - \frac{\ln[1 - \theta(1-u)]u[1 - \theta(1-u)]}{\theta u - \ln[1 - \theta(1-u)][1 - \theta(1-u)]}$
Joe	$-\ln[1 - (1-t)^\theta]$	$\left[\frac{\theta(1-t)^{\theta-1}}{[1 + (1-t)^\theta]} \right]$	$t - \frac{\ln[1 - (1-t)^\theta][1 - (1-t)^\theta]}{\theta(1-t)^{\theta-1}}$

2.9. Raegly Distrubtion

X is variable which has Raegly distribution, accordingly pdf of X,

$$y = f(x; \sigma) = \frac{x}{\sigma^2} e^{\left(\frac{-x^2}{2\sigma^2}\right)}$$

where, σ is parameter of distribution. The cumulative function of this distribution is given

$$F(x; \sigma) = 1 - e^{\frac{-x^2}{2\sigma^2}}$$

The Rayleigh distribution is a special case of the Weibull distribution [13].

3. Application

The first section, the method that has been suggested by Genest given. In this section, out of

Table 3. Non parametric estimation of Archimedean copula

Family	Gumbel	Clayton	Frank	AMH	Joe
$\hat{\theta}$	1.0298	0.059	-1.1417	0,1262	1,0517

Finally, this study consists of fitting a suitable copula to the data. The results of estimations are given table 4.

Table 4. Fitting a suitable copula the data

pairs (X, Y)	Gumbel	Clayton	Frank	AMH	Joe
	0,000275	0,004975	0,206725	0,121996	0,000001

Pearson correlation coefficient, another one measures of dependence denominated correlation. We based on Kendall Tau that has been nonparametric measures of dependence. We have seen that the doubled correlations are all positive. Namely, Kendall Tau value is positive. This study consists of estimation of Archimedean copula. Genest and Mackay (1986) simplified method and lead to estimate the parameters of Archimedean copula that focus on Kendall Tau value. Namely, Kendall Tau value is positive. This study consists of estimation of Archimedean copula. Genest and Mackay simplified method and lead to estimate the parameters of Archimedean copula that focus on Kendall Tau statistics. In this study, we generated dependent Raegly distribution $X \sim Rayl(3)$ and $Y \sim Rayl(3)$. Here $n = 200$ data were used. We calculated Kendall Tau value in such that 0.029 for (X, Y). Using this that is shown parameters of copulas calculated and results are given table 3.

4. Conclusion

In this paper, we modeled the dependence structure between $X \sim \text{Rayl}(3)$ and $Y \sim \text{Rayl}(3)$ using Archimedean copula. According to table 4, $K_n(u)$ the nonparametric estimation of $K(u)$ is calculated by using pseudo-observations, and using table 2, for Gumbel, Clayton, Frank respectively $K_G(u)$, K_C , $K_F(u)$, $K_{AMH}(u)$ and $K_J(u)$ values calculated. In table 4 $K_n(u)$ value compared $K_G(u)$, K_C , $K_F(u)$, $K_{AMH}(u)$ and $K_J(u)$. Consequently, using the square distance measure, with 0,000001 value, Joe copula gives better fits than Gumbel, Clayton, Frank and AMH.

References

- [1] Cherubini U, Luciano E. (2013), Value-at-risk trade-off and capital allocation with copulas. *Economic Notes*, vol. 30, pp. 235–256.
- [2] Fang, Hong-Bin, Kai-Tai Fang, and Samuel Kotz. (2002). The meta-elliptical distributions with given marginals. *Journal of Multivariate Analysis* vol. 82, no., pp.11-16.
- [3] Frees EW, Valdez EA. (1998). Understanding relationships using copulas. *North American Actuarial Journal*, vol. 2, pp. 1-25.
- [4] Genest C, MacKay J. (1986). The joy of copulas: bivariate distributions with uniform marginal. *The American Statistician*, vol. 40, pp. 280-283.
- [5] Genest, C., et al. (1999), A characterization of quasi-copulas. *Journal of Multivariate Analysis*, vol. 69, no.2 , pp.193-205.
- [6] Genest C, Rivest LP. (1993), Statistical inference procedures for bivariate archimedean copulas. *Journal of the American Statistical Association*, vol.88, pp.1034-1043.
- [7] Genest C, Favre AC. (2006), Everything you always wanted to know about copula modelling but were afraid to ask. *Journal of Hydrologic Engineering*, vol.12, pp. 347-368.
- [8] Genest C, Gendron M, Boudeau-Brien M. (2009), The advent of copulas in finance. *The European Journal of Finance*, vol. 15, pp. 609-618.
- [9] Mai, Jan-Frederik. (2017) *Simulating copulas: stochastic models, sampling algorithms, and applications*. World Scientific.
- [10] Malevergne Y, Sornette D . (2003), Testing the gaussian copula hypothesis for financial assets dependences. *Quantitative Finance* vol.3, pp. 231-250.
- [11] Metin A, Çalık S. (2012), Copula function and application with economic data. *Turkish Journal of Science and Technology*, vol. 7, pp. 199-204.
- [12] Karakas M.A. (2017), Dependence Structure analysis with Copula GARCH method and for data set suitable copula selection , *Natural Science and Discovery*, vol. 3, pp.13-24.
- [13] Kundu D., Ragab M. (2005), Generalized Rayleigh Distribution: Different Methods Estimation. *Computational Statistics and Data Analysis*, vol.49, pp. 187-200.
- [14] Naifar N. (2010), Modeling dependence structure with archimedean copulas and applications to the iTraxx CDS index. *Journal of Computational and Applied Mathematics*, vol. 235, pp. 2459-2466.
- [15] Nelsen R. (1999), *An Introduction to Copulas*. Springer-Verlag, New York, USA. 272p.
- [16] Oh, Dong Hwan, and Andrew J. Patton. (2017), Modeling dependence in high dimensions with factor copulas." *Journal of Business & Economic Statistics*, vol.35, no.1, pp. 139-154.
- [17] Rosenberg J, Schuermann T. (2006), A general approach to integrated risk management with Skewed, Fat-tailed Risks. *Journal of Financial Economics*, vol. 79, pp. 569-614.
- [18] Schweitzer B, Wolff EF. (1981), On nonparametric measures of dependence for random variables. *Annals of Statistics*, vol. 9, pp. 879-885.
- [19] Shih JH, Louis TA. (1995), Inferences on the association parameter in copula models for bivariate survival data. *Biometrics*, vol.51, pp. 1384-1399.
- [20] Sklar A. (1959), Fonctions de repartition a n dimensions et leurs marges. *Publications de l'Institut de Statistique de l'University de Paris*, vol.8, pp. 229-231.